

### Bernoulli RV

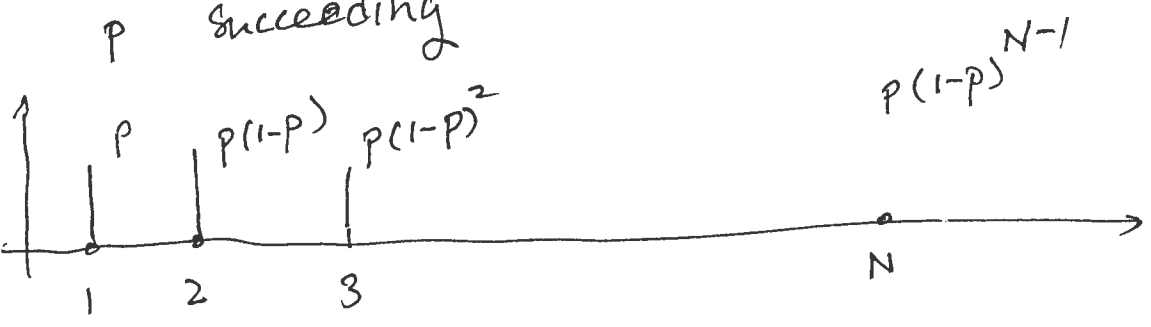
1	p
0	(1-p)

$$E[X] = p$$

$$E[(X - E[X])^2] = p(1-p)$$

### Geometric distribution

p succeeding



$$E[X] = \frac{1}{p}$$

$$VAR[X] = \frac{1-p}{p^2}$$

# Binomial

(2)

$$P(H) = p$$

flip  $N$  times

$H$  heads,  $N-H$  tails

$$\frac{N!}{H! (N-H)!} p^H (1-p)^{N-H} \rightarrow P_b(H; N, p)$$

Mean ;

$$E[X] = Np$$

$$\text{var}[X] = Np(1-p)$$

# Multinomial

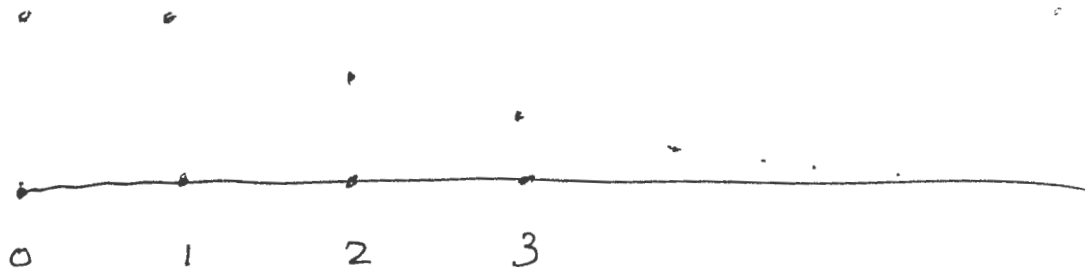
$$P(n_1, n_2, \dots, n_k; p_1, p_2, \dots)$$

$$= \frac{N!}{n_1! \cdot \dots \cdot n_k!} p_1^{n_1} \cdot \dots \cdot p_k^{n_k}$$

$$\begin{aligned} \sum_{i=1}^{\infty} i p (1-p)^{i-1} &= p \sum_{i=1}^{\infty} i (1-p)^{i-1} \\ &= p \left[ \sum_{i=1}^{\infty} (1-p)^{i-1} + \sum_{i=2}^{\infty} (1-p)^{i-1} + \dots \right] \end{aligned}$$

# Poisson distribution

$\lambda = 1$



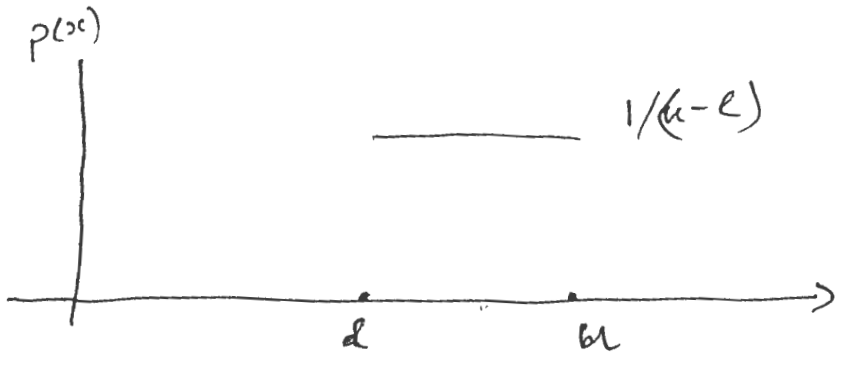
$$P(\{X=k\}) = \frac{\lambda^k e^{-\lambda}}{k!}$$

$$\sum_{k=0}^{\infty} P(\{X=k\}) = 1 = \sum_{k=0}^{\infty} \frac{\lambda^k e^{-\lambda}}{k!} = e^{-\lambda} \left[ \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} \right]$$

$$E[X] = \lambda$$

$$E[(X - E[X])^2] = \lambda$$

Uniform



# Beta - distribution

(6)

$$P_{\alpha, \beta}(x | \alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha) \Gamma(\beta)} x^{(\alpha-1)} (1-x)^{(\beta-1)}$$

$$0 \leq x \leq 1$$

$P_{\alpha, \beta}(x | \alpha, \beta)$  for  $\alpha > 1, \beta > 1$  has a single maximum  
at  $x = (\alpha - 1) / (\alpha + \beta - 2)$ .

$$E[X] = \frac{\alpha}{\alpha + \beta}$$

$$E[(X - E[X])^2] = \frac{\alpha \beta}{(\alpha + \beta)^2 (\alpha + \beta + 1)}$$

# Gamma dist

$$P_r(x|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \cdot x^{(\alpha-1)} e^{-\beta x}$$

- $x \geq 0$
- $\alpha > 0$
- $\beta > 0$

mean is  $\frac{\alpha}{\beta}$

variance is  $\frac{\alpha}{\beta^2}$

Exponential dist

$x > 0$

$$\lambda e^{-\lambda x}$$

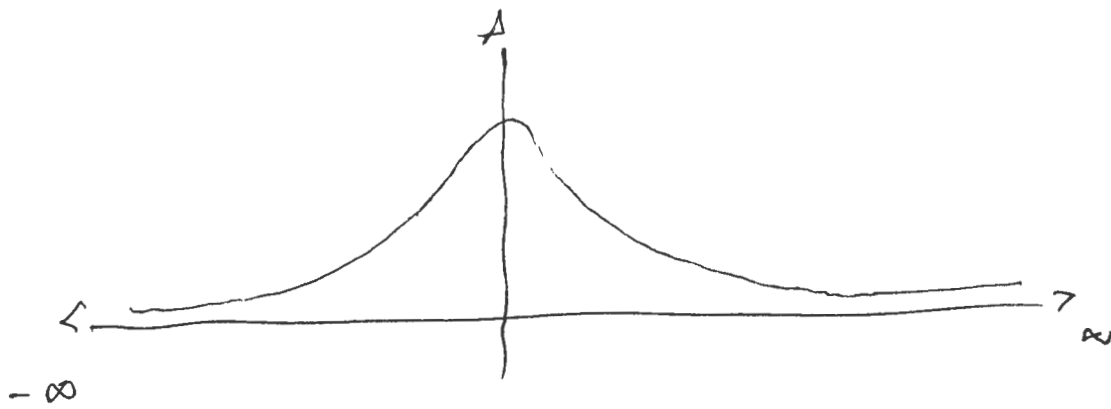
otherwise

0



# Standard normal dist

$$p(x) = \left( \frac{1}{\sqrt{2\pi}} \right) e^{-\frac{x^2}{2}}$$



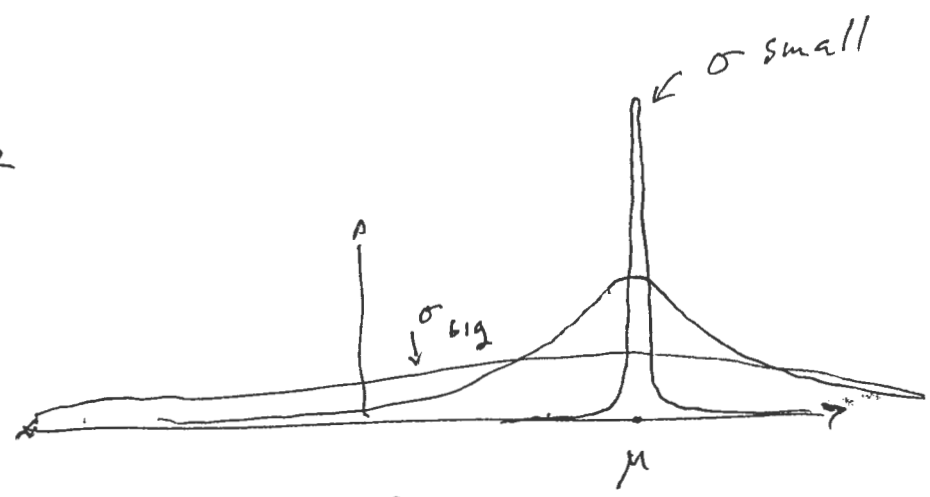
$$\int_{-\infty}^{\infty} x p(x) dx = 0 = \text{mean}$$
$$\text{Var}(x) = 1$$

Normal dist

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

mean  $\mu$

variance  $\sigma^2$



$$\int_{\mu-\sigma}^{\mu+\sigma} p(x) dx = P(\{X \in [\mu-\sigma, \mu+\sigma]\}) \approx 0.68$$

$$\mu \pm 2\sigma \approx 0.95$$

$$\mu \pm 3\sigma \approx 0.99$$

$$x = \frac{h - Np}{\sqrt{Np(1-p)}}$$

$$P(\{x \in [a, b]\}) \approx \int_a^b \frac{1}{\sqrt{2\pi}} e^{-u^2/2} du$$

$$p = 0.7$$

$$(1-p) = 0.3$$

$$\sqrt{p(1-p)} \approx 0.5$$

$$N = 1e^6$$

$$h \in \{1e3, 2e3\}$$

$$\int \frac{2e3 - 1e6 \times 0.7}{\sqrt{1e6 \times 0.3 \times 0.7}} \dots \frac{1e3 - 1e6 \times 0.7}{\sqrt{1e6 \times 0.3 \times 0.7}}$$

$$\frac{1}{\sqrt{2\pi}} \cdot e^{-u^2/2} du$$

$$N = 1e6 \quad \{ h < 1e5 \}$$

$$z = \frac{h - Np}{\sqrt{Np(1-p)}}$$

$$h < 1e5 \Rightarrow z < \frac{1e5 - 1e6 \cdot 0.7}{\sqrt{1e6 \cdot 0.7 \cdot 0.3}}$$

$$\int_{-\infty}^{(\quad)} \frac{1}{\sqrt{2\pi}} e^{-u^2/2} du$$

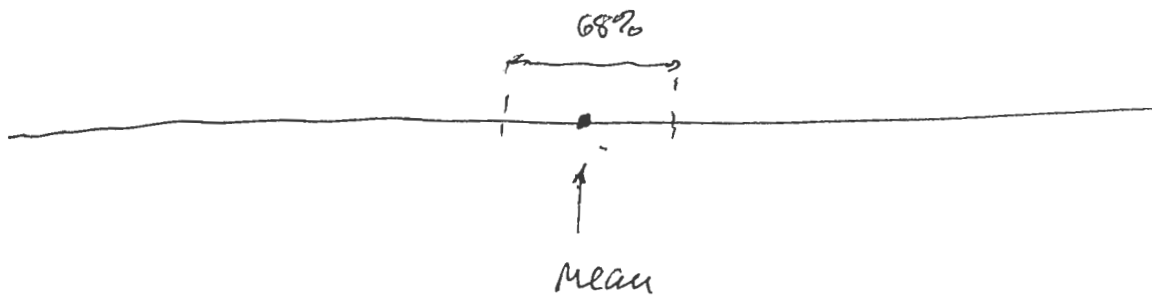
N big

P

~~h = Np~~

$$h \approx \text{Normal} \left( Np, \sqrt{Np(1-p)} \right).$$

$\uparrow$   
 mean



$$\frac{h}{N} \approx \text{Normal} \left( p, \sqrt{\frac{p(1-p)}{N}} \right)$$

$\uparrow$                        $\uparrow$   
 mean                      std



$$P(X_{k+1} = j | X_k = i)$$

$$= p_{ij}$$

$$X_0 = 1$$

$$P(X_1 = j | X_0 = 1) = p_{1j}$$

$$P(X_0 = i) = \underline{\pi}_i$$

$$P(X_1 = j) = \sum_i P(X_1 = j | X_0 = i) P(X_0 = i)$$

$$= \sum_i \pi_i p_{ij} = \underline{\pi} \underline{P}$$

$$P(X_2 = j) = \sum_k P(X_2 = j | X_1 = k) \cdot P(X_1 = k)$$

$$= \sum_k p_{kj} \left[ \sum_i \pi_i p_{ik} \right]$$

$$= \left[ \underline{\pi} \underline{P} \right] \underline{P}$$

$$\left[ \begin{array}{c} \pi \\ \underline{P}^2 \end{array} \right]$$



$$\underline{P}_s = \underline{P}_s \underline{P}$$



X is uniform  $\in [0-1]$

Y is uniform  $[0-10]$

indep.

$$Z = (Y - 5X)^3 - X^2$$

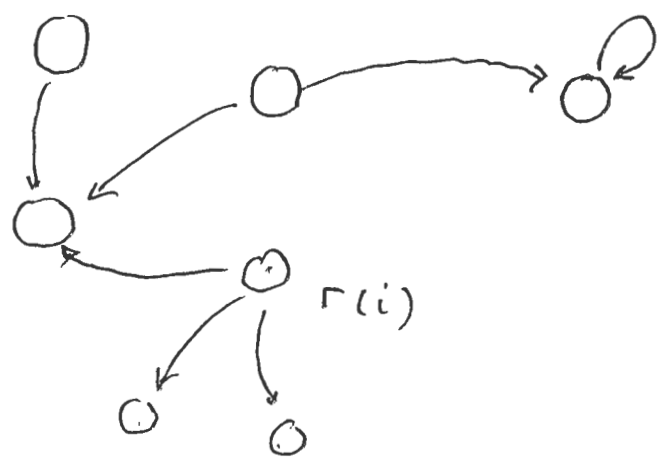
$$\text{Var}(Z) = E[Z^2] - E[Z]^2$$

$$2.76 \times 10^4$$

$$P(\{Z > 3\}) \approx 0.62$$

$$P(\{Z > 950\})$$

$$\underline{\text{std}} \approx \frac{C}{\sqrt{N}}$$



$$r(j) = \sum_{i \rightarrow j} \frac{r(i)}{|i|} = \sum P_{ij} \cdot r(i)$$

↑

$$\frac{1}{n} \cdot \frac{1}{n}$$

$$(1-\alpha) \frac{1}{|i|} + \alpha \left( \frac{1}{n} \cdot \frac{1}{n} \right)$$

↑

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# The maximum likelihood principle

$D$  - data

$p(D|\theta)$  - model

$\Theta$  - model's parameters

choose  $\hat{\theta}$  such that  $p(D|\hat{\theta})$  is a maximum

$$p(D|\theta) = \mathcal{L}(\theta)$$

N times

K heads

$\theta = p(\text{head})$

$$p(D|\theta) = \binom{N!}{k!(N-k)!} \theta^{(k)} (1-\theta)^{(N-k)}$$

$$\left( \frac{N!}{k!(N-k)!} \right) \left[ k \theta^{(k-1)} (1-\theta)^{(N-k)} - (N-k) \theta^{(k)} (1-\theta)^{(N-k-1)} \right]$$

$$k(1-\theta) - (N-k)\theta = 0$$

$$\hat{\theta} = \frac{k}{N}$$

$N$  times, stopping at first head

$$\theta = P(H).$$

$$P(D|\theta) = (1-\theta)^{(N-1)} \theta$$

$$(1-\theta)^{(N-1)} - (N-1)(1-\theta)^{(N-2)} \theta$$

$$\hat{\theta} = \frac{1}{N}$$

$$\underline{\text{Data}} = N_1, N_2$$

↑  
 $(N_1 - 1)T,$   
 followed by 1 head

$$P(N_1, N_2 | \theta) \stackrel{P(\cdot | \cdot)}{=} P(N_1 | \theta) P(N_2 | \theta)$$

$$= \left[ (1-\theta)^{N_1-1} \cdot \theta \right] \left[ (1-\theta)^{N_2-1} \cdot \theta \right]$$

$$= \left[ (1-\theta)^{(N_1+N_2-2)} \cdot \theta^2 \right]$$

$$- (N_1 + N_2 - 2) (1-\theta)^{(N_1+N_2-3)} \theta^2 + 2\theta (1-\theta)^{(N_1+N_2-2)} = 0$$

$$\theta (N_1 + N_2 - 2) = 2 \theta (1-\theta)$$

$$\hat{\theta} = \left[ \frac{2}{N_1 + N_2 - 2} \right]$$

Data  $(N_1, h_1)$ ,  $(N_2, h_2)$

$$P(D_1, D_2 | \theta) = P(D_1 | \theta) P(D_2 | \theta)$$

$$= \left[ \binom{\text{stuff}}{h_1} \theta^{h_1} (1-\theta)^{N_1-h_1} \right] \left[ \binom{\text{stuff}}{h_2} \theta^{h_2} (1-\theta)^{N_2-h_2} \right]$$

$$= \left[ \binom{\text{stuff}}{h_1+h_2} \theta^{(h_1+h_2)} (1-\theta)^{(N_1+N_2-(h_1+h_2))} \right]$$

$$\hat{\theta} = \frac{h_1+h_2}{N_1+N_2}$$

Data:

for  $n_1, n_2, n_3$

$$P(n_i | \theta) = \frac{\theta^{n_i} \cdot e^{-\theta}}{n_i!}$$

$$P(n_1, n_2, n_3 | \theta) = P(n_1 | \theta) \cdot P(n_2 | \theta) \cdot P(n_3 | \theta)$$

$$\log P(n_1, n_2, n_3 | \theta) = \log P(n_1 | \theta) + \log P(n_2 | \theta) + \log P(n_3 | \theta)$$

$$= (n_1 \log \theta - \theta) - \log(n_1!)$$

$$+ (n_2 \log \theta - \theta)$$

$$+ (n_3 \log \theta - \theta) + \text{const}$$

$$= (n_1 + n_2 + n_3) \log \theta - 3\theta + \text{const}$$

$$\hat{\theta} = \frac{\cancel{3} (n_1 + n_2 + n_3)}{\cancel{3}} = \frac{n_1 + n_2 + n_3}{3}$$



Data:

$x_1, \dots, x_N$

$$p(x|\theta) = \frac{1}{\sqrt{2\pi}\sigma} \cdot e^{-\frac{(x-\theta)^2}{2\sigma^2}}$$

$$p(x_1, x_2, \dots, x_N|\theta) = p(x_1|\theta) p(x_2|\theta) \dots p(x_N|\theta)$$

$$\log p(x_1, x_2, \dots, x_N|\theta) = \log p(x_1|\theta) + \log \dots$$

$$= -\frac{(x_1 - \theta)^2}{2\sigma^2} - \frac{(x_2 - \theta)^2}{2\sigma^2} - \dots + 'C_0$$

Care about :  $-\sum_i (x_i - \theta)^2$

$$\text{get } \hat{\theta} = \frac{\sum_i x_i}{N}$$